

Midterm Exam #2  
Tuesday, April 8, 2008

Name SOLUTION SET

Score \_\_\_\_\_

**Please do not turn this page until requested to do so.**

You may use Poor's textbook and your class notes for this exam. Make sure to show all of your work on the longer problems for partial credit. **Box your answers, and write neatly.** Backs of pages may be used for scratch work if necessary.

There are 3 problems, for a total of 110 points.

**Problem 1 (30 pts).** Let  $S_i$ ,  $1 \leq i \leq n$ , be a signal to be optimally designed,  $N_i$ ,  $1 \leq i \leq n$ , an iid  $\mathcal{N}(0,1)$  noise sequence, and  $T$  a sample selector taking values in  $\{1, 2, \dots, n\}$ . Consider the following composite hypothesis test:

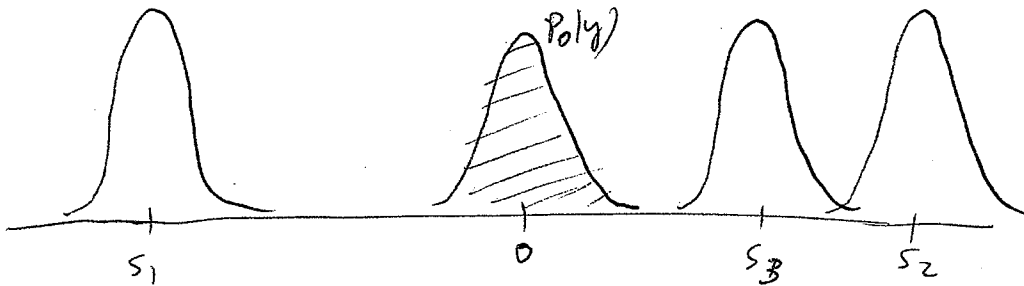
$$\begin{cases} H_0: Y = N_T, \\ H_1: Y = S_T + N_T, \quad T \in \{1, 2, \dots, n\}. \end{cases}$$

That is, the received  $Y$  is a single sample (acquired at a location  $T$  unknown to the detector) of the noise sequence or of the signal-plus-noise sequence.

Identify the signals  $S_i$ ,  $1 \leq i \leq n$ , that maximize the power of the detection test at significance level  $\alpha$ , for the worst-case  $T$ . The design is subject to the energy constraint  $\sum_{i=1}^n S_i^2 \leq n$ .

The rival pdf's are

$$\begin{cases} H_0: Y \sim \mathcal{N}(0, 1) \\ H_1: Y \sim \mathcal{N}(S_T, 1) \end{cases}$$



If all  $s_i > 0$  OR all  $s_i < 0$ , we have a UMP test:

$$\downarrow$$

$$y \underset{H_0}{\overset{H_1}{>}} \tau$$

$$\downarrow$$

$$y \underset{H_1}{\overset{H_0}{>}} \tau = Q^{-1}(\alpha)$$

The worst-case  $T$  is the one for which  $|S_T|$  is smallest.

$\Rightarrow$  "Equalizer designs"  $S_1 = S_2 = \dots = S_n = 1$

and  $S_1 = S_2 = \dots = S_n = -1$

Satisfy energy constraint with equality and maximize power of the test for worst-case  $T$ :

$$\min_T P_D = 1 - \max_T Q(|S_T - \tau|) = 1 - Q(1 - \tau)$$

Problem 2 (40 pts). Let  $N_i$ ,  $1 \leq i \leq n$ , be iid Bernoulli random variables with  $Pr[N = 1] = \theta \leq \frac{1}{2}$ . Consider the following binary hypothesis test:

$$\begin{cases} H_0: Y_i = N_i, \\ H_1: Y_i = S_i \oplus N_i, \quad 1 \leq i \leq n, \end{cases}$$

where  $\underline{S} = \{S_i, 1 \leq i \leq n\}$ , is a fixed signal, and  $\oplus$  denotes modulo-2 addition.

(a) (20 pts) Assuming equal priors and uniform costs, derive the Bayes rule for this test.

Let  $\tilde{n} = \sum_{i=1}^n 1_{\{S_i=1\}} = \text{Hamming weight of } \underline{S}. \quad \leq n$

$$I = \{i: S_i = y_i = 1\}$$

$$\tilde{n}_1 = |I| \leq \tilde{n}$$

With this notation we have

LRT: 
$$\frac{P_1(\underline{y})}{P_0(\underline{y})} = \frac{(1-\theta)^{\tilde{n}_1} \theta^{\tilde{n}-\tilde{n}_1}}{\theta^{\tilde{n}_1} (1-\theta)^{\tilde{n}-\tilde{n}_1}} = \left(\frac{1-\theta}{\theta}\right)^{2\tilde{n}_1 - \tilde{n}} \quad \begin{matrix} H_1 \\ \sum \tau = 1 \\ H_0 \end{matrix}$$

$$\Rightarrow \tilde{n}_1 \underset{H_0}{\overset{H_1}{\geq}} \frac{1}{2} \left( \tilde{n} + \frac{\ln \tau}{\ln \frac{1-\theta}{\theta}} \right) \Rightarrow \boxed{\tilde{n}_1 \underset{H_0}{\overset{H_1}{\geq}} \frac{1}{2} \tilde{n}}$$

(b) (10 pts) Determine the signal  $\underline{S}$  that minimizes the error probability above.  
of the test

$P_e = P_0 \left[ \tilde{N}_1 > \frac{1}{2} \tilde{n} \right]$  is minimized when  $\tilde{n} = n$ ,  
i.e.,  $\boxed{S_1 = S_2 = \dots = S_n = 1}$  (all-one sequence)

(c) (10 pts) Give a large-deviations upper bound on the error probability above.

$$P_e \leq e^{-n \mu^*(1/2)}$$

where  $\mu^*(a) = a \log \frac{a}{\theta} + (1-a) \log \frac{1-a}{1-\theta}$ ,  $0 < a < 1$   
(HW 4 Prob 4(a))

$$\Rightarrow \mu^*(1/2) = -\frac{1}{2} \log [4\theta(1-\theta)]$$

$$\Rightarrow \boxed{P_e \leq [4\theta(1-\theta)]^{n/2}}$$



- (b) (10 pts) Use the Central Limit Theorem to approximate  $\Pr\left[\underbrace{\sum_{i=1}^n X_i}_{=\sigma_n} > n + n^\alpha\right]$ , where  $\frac{1}{2} \leq \alpha < 1$ .

$$E[X] = \text{Var}[X] = 1 \quad \Rightarrow \begin{cases} E[\sigma_n] = n \\ \text{Var}[\sigma_n] = n \end{cases}$$

$$\Rightarrow \hat{P}_{\text{CLT},n} = Q\left(\frac{n^\alpha}{\sqrt{n}}\right) = \boxed{Q(n^{\alpha-1/2})}$$

$$\text{As } n \rightarrow \infty, \text{ we have } \hat{P}_{\text{CLT},n} \sim \frac{1}{\sqrt{2\pi n^{\alpha-1/2}}} e^{-\frac{1}{2} n^{2\alpha-1}}$$

- (c) (10 pts) Use the large-deviations method of Part (a) to derive an upper bound on the tail probability of Part (b). You may use asymptotic approximations, e.g.,  $\ln(1+\epsilon) \sim \epsilon - \frac{1}{2}\epsilon^2$  as  $\epsilon \rightarrow 0$ . Compare your result with that of Part (b).

Following the steps of Part (a), we obtain (observe  $n + n^\alpha > E[\sigma_n] = n$ )

$$\begin{aligned} \Pr\left[\underbrace{\sum_{i=1}^n X_i}_{=\sigma_n} > n + n^\alpha\right] &\leq e^{-[S(n+n^\alpha) - \mu_{\sigma_n}(s)]} \\ &= e^{-n \left[ \underbrace{s(1+n^{\alpha-1})}_{a \downarrow 1} - \log \frac{1}{1-s} \right]} \end{aligned}$$

$$\Rightarrow \Pr[\sigma_n > n + n^\alpha] \leq e^{-n \mu^*(1+n^{\alpha-1})}$$

$$\begin{aligned} \text{where } \mu^*(a) &= a - \ln a - 1 \\ &\sim \frac{1}{2}(a-1)^2 \quad \text{as } a \rightarrow 1 \end{aligned}$$

$$\mu^*(1+n^{\alpha-1}) \sim \frac{1}{2} n^{2\alpha-2}$$

$$\leq \boxed{e^{-\frac{1}{2} n^{2\alpha-1}}}$$

The CLT approximation of Part (b) correctly captures the exponent.

NOTE. The regime  $\frac{1}{2} \leq \alpha < 1$  is known as the "moderate-deviations" regime (Dembo).